

# Agenda: Information, Probability, Statistical Entropy

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- Information and probability
  - simple combinatorics.
  - Probability distributions, joint probabilities.
  - stochastic variables, correlations.
  - Statistical entropy
  - Monte Carlo simulations
- <sup>1</sup> • Partition of probability
- Phase space evolution (Eta Theorem)
- Partition functions for different degrees of freedom
- Gibbs stability criteria, equilibrium

## Reading Assignments

Weeks 4 & 5

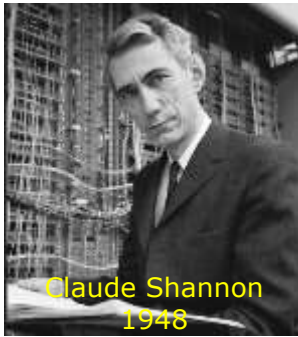
LN III.1- III.6:

Kondepudi Ch. 20  
Additional Material

McQuarrie & Simon  
Ch. 3 & 4

Math Chapters  
MC B, E

# Aspects of Information Theory



In the absence of information, probability replaces certainty.

Information theory provides probability as an objective link between randomness and certainty.

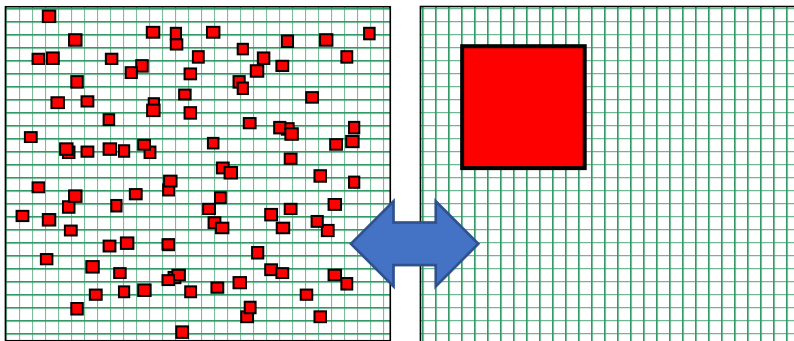
$$H = -\sum p(x) \log p(x)$$

Important theorist

Claude Shannon, "Father of Information Theory"

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"If a situation (event) is very likely to occur (high probability) → information provided with its occurrence is low." And vice versa.



Random positions  
= first guess if  
nothing known  
about particles

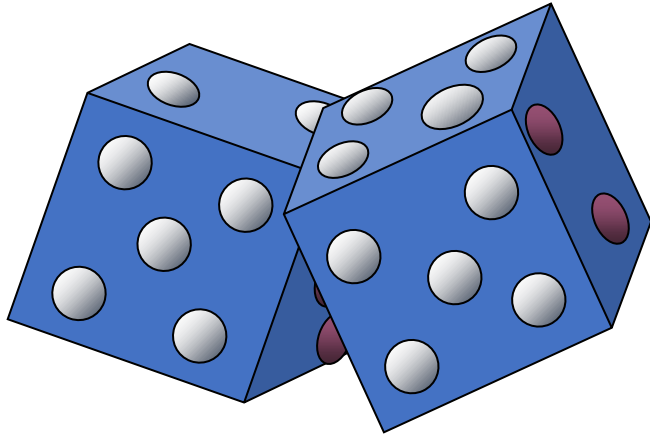
Particles cluster in  
a corner → deduce  
mutual attraction  
→ significant info

Illustration:

Two possible 2-dim micro-states for a system of 100 particles distributed differently over the available phase space.

# Simple Probability Concepts

W. Udo Schröder 2023



Probability: understood as being relative to set of many uncertain events, experiments, measurements, outcomes,

...

If instances of a type of event  $x_n$  occur in random selection, without hidden preference (bias), one can estimate

Probability  $\leftrightarrow$  combinatorics

Conduct experiment/measurement  $\rightarrow$  probability for event type  $x$

$N(x_k) = \text{frequency of events } x_k$

Probability  
**a posteriori**

$$P(x_n) = \lim_{N \rightarrow \infty} \frac{N(x_n)}{\sum_k N(x_k)}$$

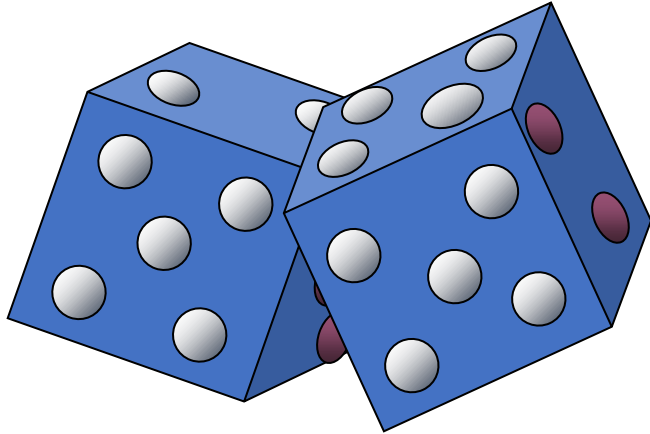
**Example:** experiment (unbiased) rolling 1 dice 1000x  $\rightarrow$

Expect (**a priori**) face with "6"  $P(6) = 1/6 = 0.1667 \rightarrow N(6) = 166$  (167)

observe (**a posteriori**) face with "6" :  $N(6) = 173 \rightarrow P(6) = 0.173 \approx 1/6$

# Probability Distributions

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Conduct many ( $M=5$ ) measurements of 1000 rolls of one die  $\rightarrow M=5$  different probabilities for event type  $x = "6"$   
 $\rightarrow$  **a posteriori** probability

$N(x_k) =$  frequency of events  $x_k$

$\{N(6)\} = \{175, 160, 155, 167, 182\}$

$$P_i(6) = \lim_{N \rightarrow \infty} \frac{N_i(6)}{\sum_{k=1}^6 N_i(k)}; \quad i = 1, \dots, M$$

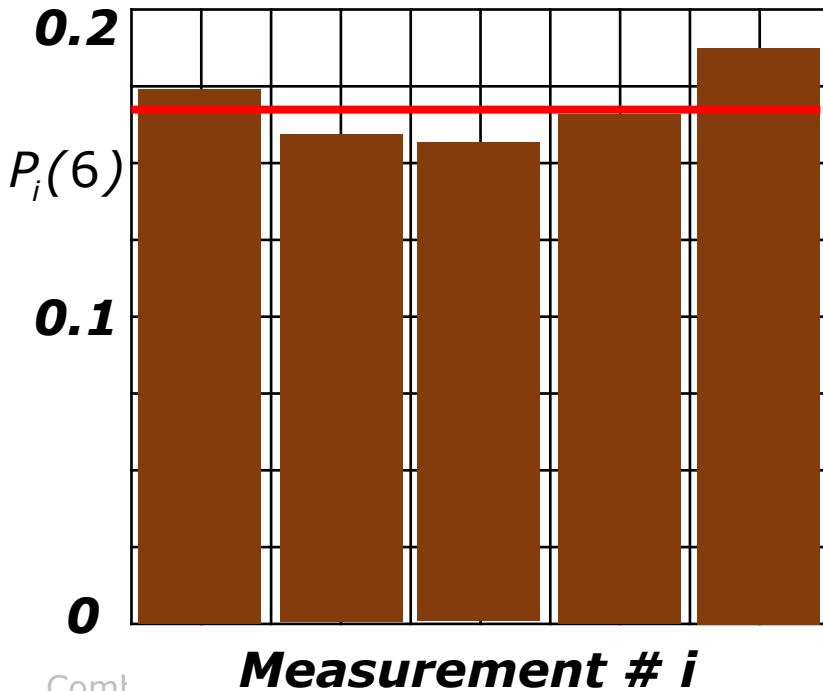
**Mean/average probability**

$$\langle P(6) \rangle_i := \sum_{i=1}^M w_i \cdot P(6)_i = \sum_{i=1}^M \frac{1}{M} \cdot P(6)_i$$

Measurements of equal quality

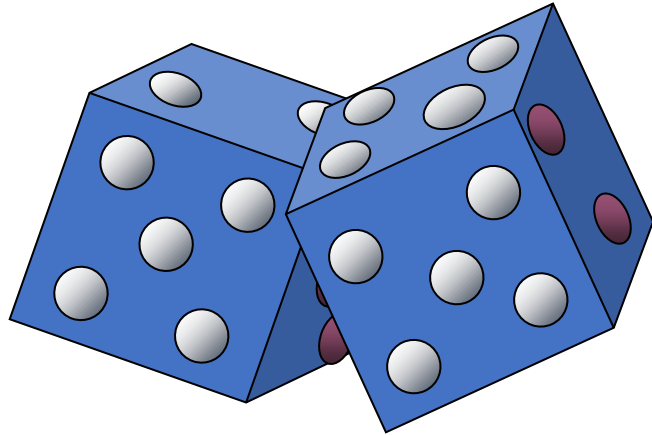
$\rightarrow$  Equal weights  $w_i = 1/M$

$\rightarrow \langle P(6) \rangle_i = 0.168$



Comt

# Probability Distributions



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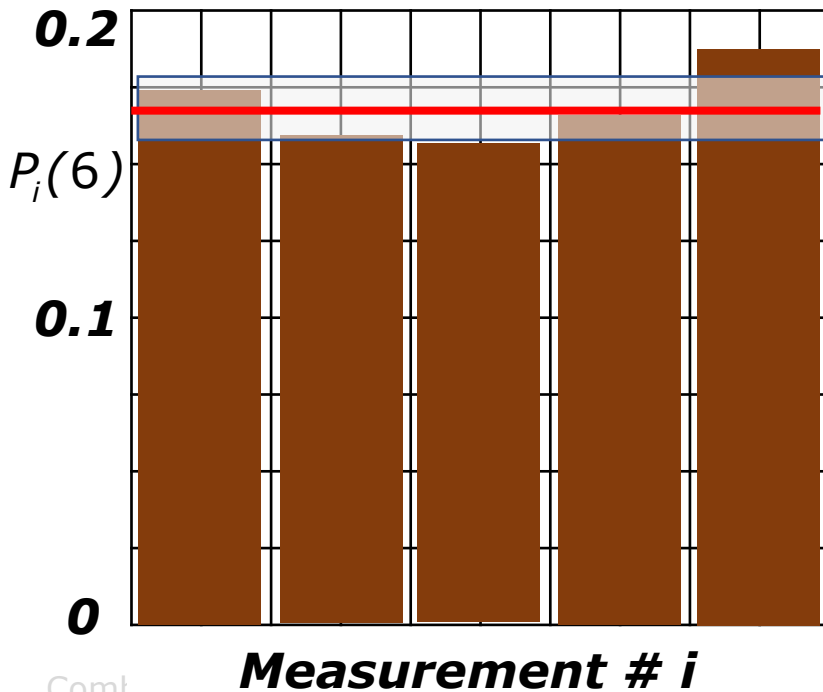
## Variance/Standard Deviation

$$\sigma_P^2(6) := \frac{1}{M(M-1)} \sum_{i=1}^M (P(6)_i - \langle P \rangle)^2$$

$$\sigma_P^2(6) = 2.4 \cdot 10^{-5} \rightarrow \sigma_P(6) = 0.005$$

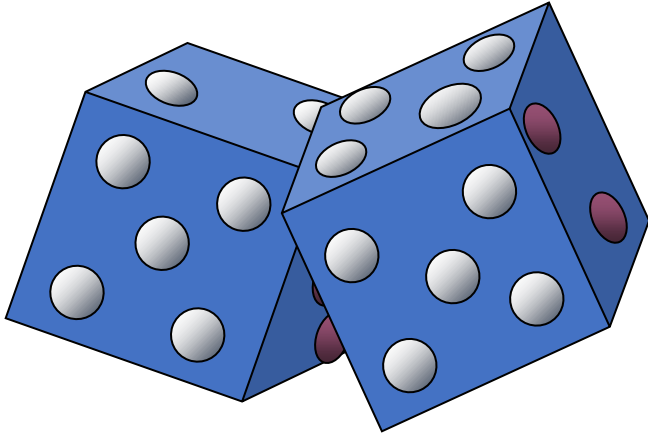
$$\rightarrow \langle P(6) \rangle_i = (0.168 \pm 0.005)$$

$$a \text{ priori} : P(6) = 1/6 = 0.167$$



# Probability Combinations

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Additive, inclusive or exclusive, cumulative, multiplicative, conditional probabilities.

Sum & product rules for disjoint (independent) probabilities

Outcome of one trial ( $\rightarrow$ Event E1) has no effect on the result of the next trial ( $\rightarrow$ Event E2). The corresponding probabilities are independent of one another and add

**A priori** probability to get a face “6” in either of two trials, the first or the second throw of a dice, equals the sum of both

$$P_{1\vee 2}(6) = P_1(6) + P_2(6) = \frac{1}{6} + \frac{1}{6} = \frac{1}{3} \quad \text{OR inclusive}$$

**Simultaneous (joint) a priori** probability to get a face “6” **in both** trials, the first **and** the second throw of a dice, equals the **product** of both

$$P_{1\wedge 2}(6 | 6) = P_1(6) \cdot P_2(6) = \frac{1}{6} \cdot \frac{1}{6} = \frac{1}{36} \quad \text{AND}$$

# Properties of *a priori* Probabilities

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- The probability for an event is  $0 \leq P \leq 1$
- The probability of an impossible event is zero,  $P = 0$ .
- The probability for one of the possible outcomes to occur is  $P = \sum_i P_i = 1$
- If two events (E1 and E2) are independent (disjoint, mutually exclusive), the probability of the sum (“OR”) event is the sum of the probabilities,  $P_{1 \vee 2} = P_1 + P_2$
- If two events are not mutually exclusive,  $P_{1 \vee 2} = P_1 + P_2 - P_{1 \wedge 2}$
- If two events (E1 and E2) are independent (not mutually exclusive), the probability for simultaneous occurrence is the product  $P_{1 \wedge 2} = P_1 \cdot P_2$
- Additional considerations for conditional (*marginal*) probabilities, example  $P\{E_1 | E_2\} := \text{Probability}\{E_1\}, \text{ given } E_2$

# Conditional Probabilities

Constraints on a set of events  $\rightarrow$  conditional probability  $P\{E|\text{condition}\}$ .

Example drawing from 3 balls (R, G, G) hidden in a box. What are *a priori* probabilities for a **sequence 1. draw, 2. draw,....** e.g.  $G_2, R_1$ ...

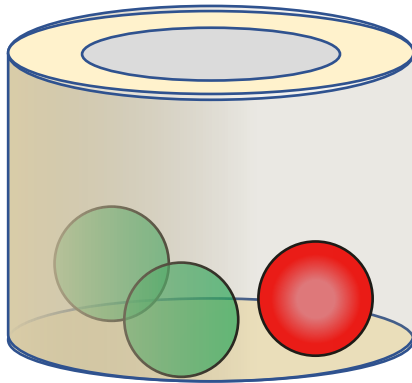
*Depends on* 1. draw returned or not! **If returned, uncorrelated draws  $\rightarrow$**

$$P(G_2 R_1) = P(G_2) \cdot P(R_1) = \frac{2}{3} \cdot \frac{1}{3} = \frac{2}{9}$$

If ball is not returned, correlated=conditional draws  $\rightarrow$

Given that *Red* was drawn first

$$P(G_2 R_1) \rightarrow P\{G_2 | R_1\} \cdot P(R_1) = 1 \cdot \frac{1}{3} = \frac{1}{3}$$



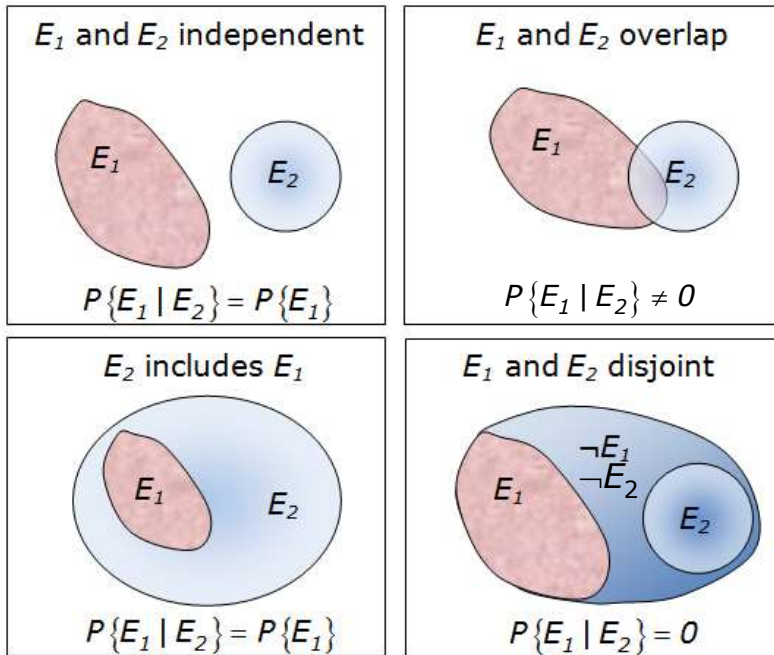
3. draw

$$P(G_3 G_2 R_1) = P\{G_3 | G_2 R_1\} \cdot P(G_2 R_1) = 1 \cdot \frac{1}{3} = \frac{1}{3}$$



# Statistical Event Domains

Possible relations are illustrated between hypothetical domains of probable events



$$P\{E_1\}, P\{E_2\}, \text{ and } P\{\neg E_1 \vee \neg E_2\}$$

Independent events  $E_1$  and  $E_2$  or if  $E_1 \subset E_2 \rightarrow E_2$  has no influence on the probability for  $E_1 \rightarrow$  **Conditional**

$$P\{E_1 | E_2\} \text{ (given } E_2 = T) = P\{E_1\}$$

If two events are mutually exclusive,

$$P\{E_1 \wedge E_2\} = 0 = P\{E_1 | E_2\}$$

If  $E_1, E_2$  not disjoint :

$$\begin{aligned} P\{E_1 \wedge E_2\} &= P\{E_1 | E_2\} \cdot P(E_2) \\ &= P\{E_2 | E_1\} \cdot P(E_1) \end{aligned}$$

Fractional probability for  $E_2 = T$  (if  $E_1 = T$ ) / Total  $P(E_2)$

## Bayes' Theorem

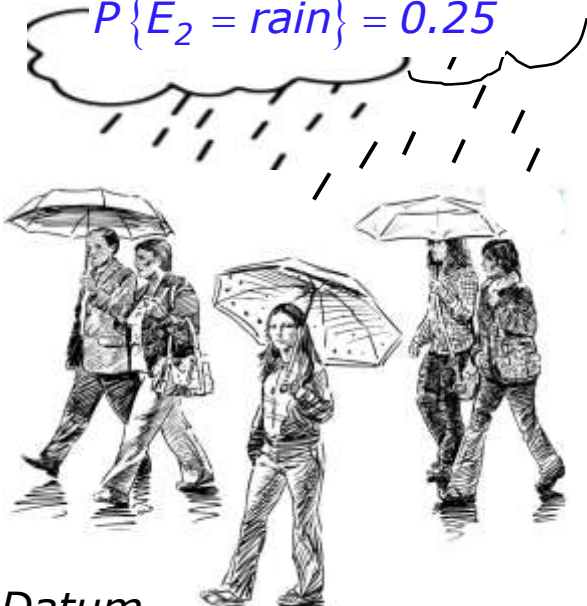
$$P\{E_1 | E_2\} = P(E_1) \cdot \frac{P\{E_2 | E_1\}}{P(E_2)}$$

Prior probability  $P(E_1) =$  suspicion, guess

# Bayes' Theorem: Rain or Sun ?

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$P\{E_2 = \text{rain}\} = 0.25$



*Datum*  
 $P\{E_1 \mid \text{given } E_2 = \text{rain}\} = 0.8$


$E_2 = \text{partly cloudy}$   
 $P\{E_2\} = 0.42$

Event  $E_1 =$   
*people carry*  
*umbrellas*  
 $P\{E_1\} = 0.20$

$E_2 = \text{other}$   
*reasons*

$P\{E_1 \mid E_2\} = 0.1$

$P\{E_2 = \text{sun}\} = 0.33$



*Datum*  
 $P\{E_1 \mid \text{given } E_2 = \text{sun}\} = 0.1$

*Prob for  $E_2 = \text{rain}$  when observing umbrellas*

$$P\{E_2 = \text{rain} \mid E_1 - \text{umbr.}\} = P(E_2 = r) \cdot \frac{P\{E_1 = u \mid E_2 = r\}}{P(E_1 = u)} = 0.25 \cdot \frac{0.8}{0.2} = 1$$

# Discrete Multivariate Probability Distributions

**Bias Test:** Randomly draw colored dies ( $i=1-5$ ) out of bag. Roll each die many times and record face frequencies ( $j=1,\dots,6$ ). Normalize to total # rolls.

→  $\dim\{\mathbf{p}_{ij}\}$  domain =  $5 \times 6 = 30$  → **Data must fulfill independent constraints:**

$$\text{Color } u_i = \sum_{j=1}^6 p_{ij} \rightarrow \sum_{i=1}^5 u_i = 1 \quad \text{Face } v_j = \sum_{i=1}^5 p_{ij} \rightarrow \sum_{j=1}^6 v_j = 1$$

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Color $i$	Face $j$						$u_i = \sum_{j=1}^6 p_{ij}$
	1	2	3	4	5	6	
red							$u_1 = 0.10$
blue				$p_{24}$			$u_2 = 0.30$
green							$u_3 = 0.20$
white							$u_4 = 0.25$
black							$u_5 = 0.15$
$v_j = \sum_{i=1}^5 p_{ij}$	$v_1$ 0.131	$v_2$ 0.148	$v_3$ 0.163	$v_4$ 0.162	$v_5$ 0.200	$v_6$ 0.196	Data

After Dill & Bromberg

# Discrete Multivariate Probabilities






Compare a priori with posteriori probabilities to find bias. Example: blue die with face "4" from overlap (simultaneous) probability domains.

A priori :  $P(\text{blue}, \text{any } \#) = 1/5$ ;  $P(\text{any}, 4) = 1/6 \rightarrow P(\text{blue}) \wedge P(4) = 1/30 = 0.033$

A posteriori :  $P(\text{blue}) \wedge P(4) = u_2 \cdot v_4 = 0.3 \cdot 0.162 = 0.049$

But  $0.049/0.033 \approx 1.5 = u_2/0.2$

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		$j \rightarrow$						
	$i \downarrow$	1	2	3	4	5	6	$u_i = \sum_{j=1}^6 p_{ij}$
	red	0.016	0.018	0.018	0.017	0.015	0.016	$u_1 = 0.10$
	blue	0.025	0.035	0.045	0.055	0.065	0.075	$u_2 = 0.30$
	green	0.030	0.028	0.032	0.030	0.050	0.030	$u_3 = 0.20$
	white	0.040	0.042	0.038	0.040	0.040	0.050	$u_4 = 0.25$
	black	0.020	0.025	0.030	0.020	0.030	0.025	$u_5 = 0.15$
	$v_j = \sum_{i=1}^5 p_{ij}$	$v_1$	$v_2$	$v_3$	$v_4$	$v_5$	$v_6$	
		0.131	0.148	0.163	0.162	0.200	0.196	

Data from Dill & Bromberg

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# End

## Probability & Combinatorics