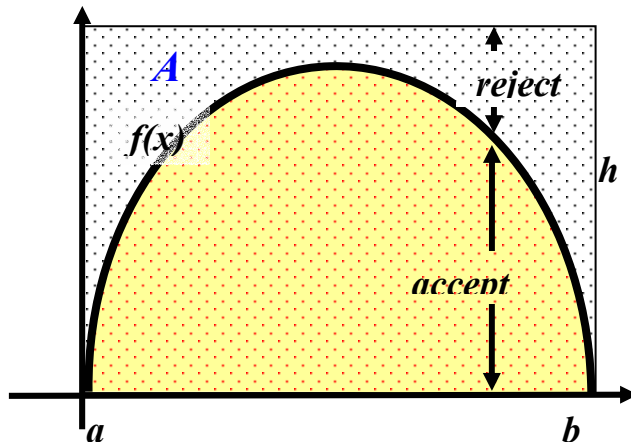


Monte Carlo Random Sampling



The [Monte Carlo](#) random sampling method generates what can be called "*artificial data*" of a prescribed form, represented, e.g., by a functional relation $y(x)$ between independent (x) and dependent (y) variables. The simplest Monte Carlo simulation method is called the *Rejection Method*. It is used in

applications with normally distributed data sets in the *MATHCAD* code [mathcad\MonteCarlo-01i.mcd](#). The method is very useful in the numerical evaluation of multi-dimensional integrals of possibly complicated functions of several variables, which may not have analytical primitive functions.

In the following, this method is illustrated with the simple case of a numerical integration of a function $f(x)$ of one variable x over the interval $[a, b]$. As shown in the sketch, one draws a graph of the function $f(x)$ and encloses that part of the function to be integrated within a (not necessarily, but *preferably the smallest*) rectangle of area A , in the example $A = h \cdot (b - a)$. Then, a large number of pairs of real numbers $\{x_i, y_i\}$ are chosen randomly within the area A . Each pair $\{x_i, y_i\}$ is tested as to whether or not $y_i \leq f(x_i)$, i.e., whether or not the point $\{x_i, y_i\}$ lies within the area *below* the curve $f(x)$ (see figure above). If this is the case, the point is accepted for calculating the integral $\int_a^b dx f(x)$, otherwise, the point is rejected. Counting the number N_{acc} of accepted points out of the total number N_{tot} of random points drawn, one has a measure of the

integral, since the area under the curve scales to the total area like the ratio N_{acc}/N_{tot}

$$\int_a^b dx f(x) = \frac{N_{acc}}{N_{tot}} \cdot A \quad (1)$$

This is a plausible example of the general Monte Carlo integration rule

$$\int_a^b dx f(x) \approx \langle f \rangle (b-a) \pm \sigma_f (b-a) \quad (2)$$

with the statistical error involving the standard deviation

$$\sigma_f \approx \sqrt{\frac{\langle f^2 \rangle - \langle f \rangle^2}{N}} \quad (3)$$

The averages and errors (see also: [Prob-Moments.doc](#)) are defined with respect to the functional values at the x coordinates of the N sampling points $\{x_i, y_i\}$:

$$\langle f \rangle = \frac{1}{N} \sum_{i=1}^N f(x_i) \quad \text{and} \quad \langle f^2 \rangle = \frac{1}{N} \sum_{i=1}^N f^2(x_i) \quad (4)$$

The \pm term in Equ. 2 with the standard deviation of Equ. 3 represents only an error *estimate* for the value of the integral obtained with N points. For a more accurate error determination of the Monte Carlo integration method, books on numerical analysis should be consulted.

It is straight-forward to check the consistency of the example with the general method defined in Equ. 2. Realizing that, for the example, $\langle f \rangle = h \cdot (N_{acc} / N_{tot})$, and inserting this into Equ. 2, one obtains the earlier result of Equ.1. At the same time, it is obvious from Equ. 2 that

$$f(x) \approx \langle f(x) \rangle \pm \sigma_f \quad (5)$$

with

$$\langle f(x) \rangle = \frac{N_{acc}(x)}{N_{tot}(x)} \cdot h \quad (6)$$

Here, the sampling is done at a particular abscissa value x , i.e., the set $\{(x_i, y_i) | x_i = x\}$ is chosen and the number N_{acc} of acceptable points is counted out of a total of N_{tot} draws. Then, the number of accepted points is proportional to the average of the function f at the argument x , namely

$$N_{acc}(x) = \langle f(x) \rangle \frac{N_{tot}(x)}{h} \quad (7)$$

It is approximately equal to the actual value of the function f at argument x ,

$$N_{acc}(x) = \langle f(x) \rangle \frac{N_{tot}(x)}{h} \pm \sigma_f \frac{N_{tot}(x)}{h} \quad (8)$$

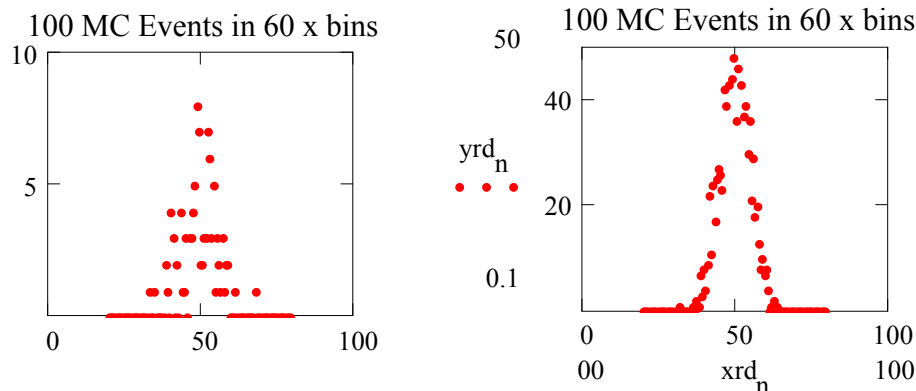
where the second term indicates again the estimated uncertainty.

It follows from the above discussion that *randomly chosen pairs $\{x_i, y_i\}$ with $y_i \leq f(x_i)$ have a probability distribution given by the function $f(x)$,*

$$P(x) = \frac{N_{acc}(x)}{N_{tot}(x)} \propto f(x) \quad (9)$$

For example, there are many pairs with x -values where the magnitude of f is large, and very few pairs represent the function where it is small. It also does not matter, in which sequence the various arguments x are sampled. So, normally the pairs $\{x_i, y_i\}$ are chosen at random with respect to both x and y . Since $P(x)$ is given by the *fraction* of the total number of points drawn, it is already *properly normalized* to unity. This is the essence of Monte Carlo random sampling.

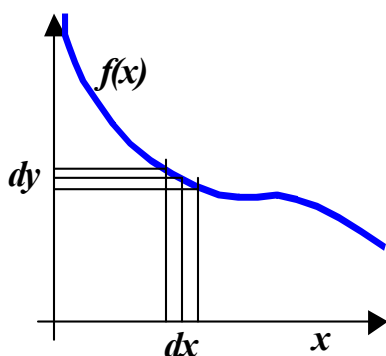
The principle is applied to a [normal](#) (*Gaussian distributed*) sample of x values with an average value of $\langle x \rangle = x_0 = 50$ and a standard deviation of $\sigma_x = 5$. The two figures above represent the same Gaussian distribution, on the left a very small sample of only 100 events in ± 30 bins about the average, and on the right 1000 events over the same range of x values. In either figure, the frequency



$P(x)$ of finding the value x in the sample of accepted points (cf. Equ. 9) is plotted vs. x . It is obvious that a small sample gives only a rough idea of the general shape of the function to be simulated numerically, here a Gaussian, with significant *statistical errors*. Increasing the sample size by a factor 10, to 1000, leads to a much

more accurate representation, as illustrated by the figure on the right.

The rejection method discussed above is universally applicable. However, depending on the functional relationship to be simulated,



a large fraction of the randomly drawn pairs $\{x_i, y_i\}$ may be rejected and, hence, useless. If this is the case, the method is not very effective and the computing time needed to generate a large statistical sample may become very long.

In special cases, the alternative **Transformation Method** may be faster. This latter method makes use of the bunching

bunching effected by an appropriate reflection of a randomly distributed set $\{x_i\}$ of x values at a function $f(x)$. As illustrated in the sketch, if the x -values are uniformly distributed, the values $y = f(x)$ are compressed, as given by the slope dy/dx . The smaller this slope, the stronger the bunching of the corresponding y -values:

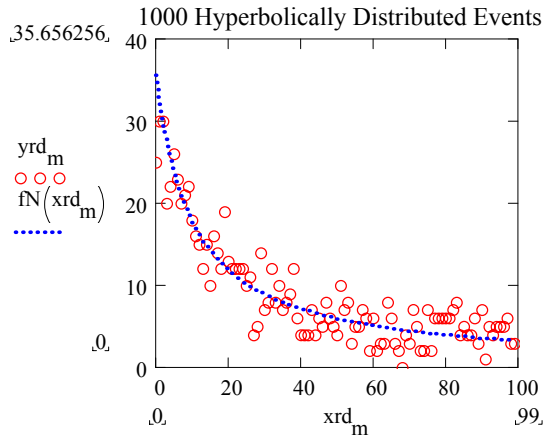
$$dy = f'(x) \cdot dx \quad (10)$$

where f' is the derivative of f . If the x -values are random, i.e., $P(x) = \text{const.}$, then the y -values are distributed according to

$$P(y) = P(x) \cdot \left(\frac{dy}{dx}\right)^{-1} \propto (f'(x))^{-1} \quad (11)$$

In other words, according to Equ. 11, choosing a random sample $\{x_i\}$, the values of the set $\{y_i = f(x_i)\}$ simulate the function $(f'(x))^{-1}$.

Suppose that one wishes to simulate a probability distribution $P(x) = f(x)$ instead of $(f'(x))^{-1}$. If it is possible to obtain a function $g(x)$, such that



$$(g'(x))^{-1} = f(x) \quad (12)$$

i.e., if it is possible to find the primitive function (integral) of $[f(x)]^{-1}$,

$$g(x) = \int dx' \frac{1}{f(x')} \quad (13)$$

then the set $\{y_i = g(x_i)\}$ built upon the random set $\{x_i\}$ has the desired probability distribution $P(x) = f(x)$.

Consider the example of a hyperbolic probability distribution

$$P(x) = f(x) = \frac{a}{b+x} \quad (14)$$

to be simulated. Obviously, the inverse of this function can easily be integrated, yielding

$$g(x) = \int dx \left[\frac{b+x}{a} \right] = \frac{b}{a}x + \frac{1}{a}x^2 \quad (15)$$

Now, choose a random set $\{x_i\}$ which can be done with the code [mathcad\MonteCarlo-01i.mcd](#) and construct the set $\{y_i = g(x_i)\}$,

also done in this code for the above example. The resulting frequency distribution $\{y_i\}$ is plotted as open circles $\{yrd_m\}$ in the figure. The analytical function $f(x)$ defined in Equ. 14 is indicated by the dashed curve (fN). Obviously, the random points outline the function well, but show a statistical scatter given by the number of events drawn.

Brief discussions of Monte Carlo Methods can be found in

R.H. Landau and M.J. Paez, *Computational Physics*, Wiley Interscience, New York, 1997

N. J. Giordano, *Computational Physics*, Prentice Hall, Upper Saddle River, 1997