| UNIVERSITY OF |  |
| :--- | :--- | :--- |
| ROCHESTER | DEPARTMENT OF CHEMISTRY <br> Monecarlo <br> W. Udo Schrider |

## Monte Carlo Random Sampling



The Monte Carlo random sampling method generates what can be called "artificial data" of a prescribed form, represented, e.g., by a functional relation $y(x)$ between independent $(x)$ and dependent $(y)$ variables. The simplest Monte Carlo simulation method is called the $\boldsymbol{R e}$ jection Method. It is used in applications with normally distributed data sets in the MATHCAD code MonteCarlo.mcd. The method is very useful in the numerical evaluation of multi-dimensional integrals of possibly complicated functions of several variables, which may not have analytical primitive functions.

In the following, this method is illustrated with the simple case of a numerical integration of a function $f(x)$ of one variable $x$ over the interval $[a, b]$. As shown in the sketch, one draws a graph of the function $f(x)$ and encloses that part of the function to be integrated within a (not necessarily, but preferably the smallest) rectangle of area $\boldsymbol{A}$, in the example $\boldsymbol{A}=\boldsymbol{h} \boldsymbol{r}(\boldsymbol{b}-\boldsymbol{a})$. Then, a large number of pairs of real numbers $\left\{x_{i}, y_{i}\right\}$ are chosen randomly within the area $\boldsymbol{A}$. Each pair $\left\{x_{i}, y_{i}\right\}$ is tested as to whether or not $y_{i} \leq f\left(x_{i}\right)$, i.e., whether or not the point $\left\{x_{i}, y_{i}\right\}$ lies within the area below the curve $\boldsymbol{f}(\boldsymbol{x})$ (see figure above). If this is the case, the point is accepted for calculating the integral $\int_{a}^{b} d x f(x), \quad$ otherwise, the point is
rejected. Counting the number $N_{a c c}$ of accepted points out of the total number $N_{\text {tot }}$ of random points drawn, one has a measure of the integral, since the area under the curve scales to the total area like the ratio $N_{a c c} / N_{\text {tot }}$

$$
\begin{equation*}
\int_{a}^{b} d x f(x)=\frac{N_{a c c}}{N_{t o t}} \cdot A \tag{1}
\end{equation*}
$$

This is a plausible example of the general Monte Carlo integration rule

$$
\begin{equation*}
\int_{a}^{b} d x f(x) \approx\langle f\rangle(b-a) \pm \sigma_{f}(b-a) \tag{2}
\end{equation*}
$$

with the statistical error involving the standard deviation

$$
\begin{equation*}
\sigma_{f} \approx \sqrt{\frac{\left\langle f^{2}\right\rangle-\langle f\rangle^{2}}{N}} \tag{3}
\end{equation*}
$$

The averages and errors (see also: Moments.doc) are defined with respect to the functional values at the $x$ coordinates of the $N$ sampling points $\left\{x_{i}, y_{i}\right\}$ :

$$
\begin{equation*}
\langle f\rangle=\frac{1}{N} \sum_{i=1}^{N} f\left(x_{i}\right) \text { and }\left\langle f^{2}\right\rangle=\frac{1}{N} \sum_{i=1}^{N} f^{2}\left(x_{i}\right) \tag{4}
\end{equation*}
$$

The $\approx$ term in Equ. 2 with the standard deviation of Equ. 3 represents only an error estimate for the value of the integral obtained with $N$ points. For a more accurate error determination of the Mon-
te Carlo integration method, books on numerical analysis should be consulted.

It is straight-forward to check the consistency of the example with the general method defined in Equ. 2. Realizing that, for the example, $\langle f\rangle=h \cdot\left(N_{a c c} / N_{t o t}\right)$, and inserting this into Equ. 2, one obtains the earlier result of Equ.1. At the same time, it is obvious from Equ. 2 that

$$
\begin{equation*}
f(x) \approx\langle f(x)\rangle \pm \sigma_{f} \tag{5}
\end{equation*}
$$

with

$$
\begin{equation*}
\langle f(x)\rangle=\frac{N_{a c c}(x)}{N_{t o t}(x)} \cdot h \tag{6}
\end{equation*}
$$

Here, the sampling is done at a particular abscissa value $x$, i.e., the set $\left\{\left(x_{i}, y_{i}\right) \mid x_{i}=x\right\}$ is chosen and the number $N_{a c c}$ of acceptable points is counted out of a total of $N_{\text {tot }}$ draws. Then, the number of accepted points is proportional to the average of the function $f$ at the argument $x$, namely

$$
\begin{equation*}
N_{a c c}(x)=\langle f(x)\rangle \frac{N_{t o t}(x)}{h} \tag{7}
\end{equation*}
$$

It is approximately equal to the actual value of the function $f$ at argument $x$,

$$
\begin{equation*}
N_{a c c}(x)=\langle f(x)\rangle \frac{N_{t o t}(x)}{h} \pm \sigma_{f} \frac{N_{t o t}(x)}{h} \tag{8}
\end{equation*}
$$

where the second term indicates again the estimated uncertainty.

It follows from the above discussion that randomly chosen pairs $\left\{x_{i}, y_{i}\right\}$ with $y_{i}=f\left(x_{i}\right)$ have a probability distribution given by the function $f(x)$,

$$
\begin{equation*}
P(x)=\frac{N_{\text {acc }}(x)}{N_{\text {tot }}(x)} \propto f(x) \tag{9}
\end{equation*}
$$

For example, there are many pairs with $x$-values where the magnitude of $f$ is large, and very few pairs represent the function where it is small. It also does not matter, in which sequence the various arguments $x$ are sampled. So, normally the pairs $\left\{x_{i}, y_{i}\right\}$ are chosen at random with respect to both $x$ and $y$. Since $P(x)$ is given by the fraction of the total number of points drawn, it is already properly normalized to unity. This is the essence of Monte Carlo random sampling.

The principle is applied to a normal (Gaussian distributed) sample


of $x$ values with an average value of $\langle x\rangle=x_{0}=50$ and a standard deviation of $\sigma_{x}=5$. The two figures above represent the same Gaussian distribution, on the left a very small sample of only 100 events in $\sim 30$ bins about the average, and on the right 1000 events
over the same range of $x$ values. In either figure, the frequency $P(x)$ of finding the value $x$ in the sample of accepted points (cf. Equ. 9) is plotted vs. $x$. It is obvious that a small sample gives only a rough idea of the general shape of the function to be simulated numerically, here a Gaussian, with significant statistical errors. Increasing the sample size by a factor 10 , to 1000 , leads to a much more accurate representation, as illustrated by the figure on the right.


The rejection method discussed above is universally applicable. However, depending on the functional relationship to be simulated, a large fraction of the randomly drawn pairs $\left\{x_{i}, y_{i}\right\}$ may be rejected and, hence, useless. If this is the case, the method is not very effective and the computing time needed to generate a large statistical sample may become very long. In special cases, the alternative Transformation Method may be faster. This latter method makes use of the bunching effected by an appropriate reflection of a randomly distributed set $\left\{x_{i}\right\}$ of $x$ values at a function $f(x)$. As illustrated in the sketch, if the $x$-values are uniformly distributed, the values $y=f(x)$ are compressed, as given by the slope $d y / d x$. The smaller this slope, the stronger the bunching of the corresponding $y$-values:

$$
\begin{equation*}
d y=f^{\prime}(x) \cdot d x \tag{10}
\end{equation*}
$$

where $f^{\prime}$ is the derivative of $f$. If the $x$-values are random, i.e., $P(x)=$ const., then the $y$-values are distributed according to

$$
\begin{equation*}
P(y)=P(x) \cdot\left(\frac{d y}{d x}\right)^{-1} \propto\left(f^{\prime}(x)\right)^{-1} \tag{11}
\end{equation*}
$$

In other words, according to Equ. 11, choosing a random sample $\left\{x_{i}\right\}$, the values of the set $\left\{y_{i}=f\left(x_{i}\right)\right\}$ simulate the function $\left(f^{\prime}(x)\right)^{-1}$.

Suppose that one wishes to simulate a probability distribution $P(x)$ $=f(x)$ instead of $\left(f^{\prime}(x)\right)^{-1}$. If it is possible to obtain a function $g(x)$, such that

$$
\begin{equation*}
\left(g^{\prime}(x)\right)^{-1}=f(x) \tag{12}
\end{equation*}
$$

i.e., if it is possible to find the primitive function (integral) of $[f(x)]^{-1}$,

$$
\begin{equation*}
g(x)=\int^{x} d x^{\prime} \frac{1}{f\left(x^{\prime}\right)} \tag{13}
\end{equation*}
$$

then the set $\left\{y_{i}=g\left(x_{i}\right)\right\}$ built upon the random set $\left\{x_{i}\right\}$ has the desired probability distribution $P(x)=f(x)$.

Consider the example of a hyperbolical probability distribution

$$
\begin{equation*}
P(x)=f(x)=\frac{a}{b+x} \tag{14}
\end{equation*}
$$

to be simulated. Obviously, the inverse of this function can easily be integrated, yielding

$$
\begin{equation*}
g(x)=\int^{x} d x\left[\frac{b+x}{a}\right]=\frac{b}{a} x+\frac{1}{a} x^{2} \tag{15}
\end{equation*}
$$

Now, choose a random set $\left\{x_{i}\right\}$ which can be done with the code MonteCarlo.mcd and construct the set $\left\{y_{i}=g\left(x_{i}\right)\right\}$, also done in this code for the above example. The resulting frequency distribution $\left\{y_{i}\right\}$ is plotted as open circles $\left\{\operatorname{yrd}_{m}\right\}$ in the figure. The analytical function $f(x)$ defined in Equ. 14 is indicated by the dashed curve ( $f N$ ). Obviously, the random points outline the function well, but show a statistical scatter given by the number of events drawn.

Brief discussions of Monte Carlo Methods can be found in
R.H. Landau and M.J. Paez, Computational Physics, Wiley Interscience, New York, 1997
N. J. Giordano, Computational Physics, Prentice Hall, Upper Saddle River, 1997

